We consider two-player games played over finite state spaces for an infinite number of rounds. At each state, the players simultaneously choose moves; the moves determine a successor state. It is often advantageous for players to choose probability distributions over moves, rather than single moves. Given a goal (e.g., “reach a target state”), the question of winning is thus a probabilistic one: “what is the maximal probability of winning from a given state?”. On these game structures, two fundamental notions are those of equivalences and metrics. Given a set of winning conditions, two states are equivalent if the players can win the same games with the same probability from both states. Metrics provide a bound on the difference in the probabilities of winning across states, capturing a quantitative notion of state “similarity”.

We introduce equivalences and metrics for two-player game structures, and we show that they characterize the difference in probability of winning games whose goals are expressed in the quantitative μ-calculus. The quantitative μ-calculus can express a large set of goals, including reachability, safety, and ω-regular properties. Thus, we claim that our relations and metrics provide the canonical extensions to games, of the classical notion of bisimulation for transition systems. We develop our results both for equivalences and metrics, which generalize bisimulation, and for asymmetrical versions, which generalize simulation.

1. Introduction

We consider two-player games played for an infinite number of rounds over finite state spaces. At each round, the players simultaneously and independently select moves; the moves then determine a probability distribution over successor states. These games, known variously as stochastic games [24] or concurrent games [3, 1, 5], generalize many common structures in computer science, from transition systems, to Markov chains [12] and Markov decision processes [6]. The games are turn-based if, at each state, at most one of the players has a choice of moves, and deterministic if the successor state is uniquely determined by the current state, and by the moves chosen by the players.

It is well-known that in such games with simultaneous moves it is often advantageous for the players to randomize their moves, so that at each round, they play not a single “pure” move, but rather, a probability distribution over the available moves. These probability distributions over moves, called mixed moves [20], lead to various notions of equilibria [29, 20], such as the equilibrium result expressed by the minimax theorem [29]. Intuitively, the benefit of playing mixed, rather than pure, moves lies in preventing the adversary from tailoring a response to the individual move played. Even for simple reachability games, the use of mixed moves may allow players to win, with probability 1, games that they would lose (i.e., win with probability 0) if restricted to playing pure moves [3]. With mixed moves, the question of winning a game with respect to a goal is thus a probabilistic one: what is the maximal probability a player can be guaranteed of winning, regardless of how the other player plays? This probability is known, in brief, as the winning probability.

*This research was sponsored in part by the grants NSF-CCF-0427202, NSF-CCF-0546170, and NSF-CCR-0132780.

In LICS 07: Proceedings of the 22nd Annual IEEE Symposium on Logic in Computer Science, 2007
In structures ranging from transition systems to Markov decision processes and games, a fundamental question is the one of equivalence of states. Given a suitably large class $\Phi$ of properties, containing all properties of interest to the modeler, two states are equivalent if the same properties hold in both states. For a property $\phi$, denote the value of $\phi$ at $s$ by $\phi(s)$: in the case of games, this might represent the maximal probability of a player winning with respect to a goal expressed by $\phi$. Two states $s$ and $t$ are equivalent if $\phi(s) = \phi(t)$ for all $\phi \in \Phi$. For (finite-branching) transition systems, and for the class of properties $\Phi$ expressible in the $\mu$-calculus [14], state equivalence is captured by bisimulation [19]; for Markov decision processes, it is captured by probabilistic bisimulation [22]. For quantitative properties, a notion related to equivalence is that of a metric: probabilistic bisimulation [22]. For quantitative properties, the metric distance of two states having metric distance 0). The metrics and relations represent the canonical extension of simulation. In the definition of bisimulation for transition systems, for every pair $s$, $t$ of bisimilar states, we require that if $s$ can go to a state $s'$, then $t$ should be able to go to $t'$, such that $s'$ and $t'$ are again bisimilar (we also ask that $s$, $t$ have an equivalent predicate valuation). This definition has been extended to Markov decision processes by requiring that for every mixed move from $s$, there is a mixed move from $t$, such that the moves induce probability distributions over successor states that are equivalent modulo the underlying bisimulation [22, 21]. Unfortunately, the generalization of this appealing definition to games fails. It turns out, as we prove in this paper, that requiring players to be able to replicate probability distributions over successors (modulo the underlying equivalence) leads to an equivalence that is too fine, and that may fail to relate states at which the same quantitative $\mu$-calculus formulas hold. We show that phrasing the definition in terms of distributions over successor states is the wrong approach for games; rather, the definition should be phrased in terms of expectations of certain metric-bounded quantities.

Our starting point is a closer look at the definition of metrics for Markov decision processes. We observe that we can manipulate the definition of metrics given in [28], obtaining an alternative form, which we call the $a$ priori form, in contrast with the original form of [28], which we call the $a$ posteriori form. Informally, the $a$ posteriori form is the traditional definition, phrased in terms of similarity of probability distributions; the $a$ priori form is instead phrased in terms of expectations. We show that, while on Markov decision processes these two forms coincide, this is not the case for games; moreover, we show that it is the $a$ priori form that provides the canonical metrics for games.

We introduce metrics and equivalence relations for concurrent games, with respect to the class of properties $\Phi$ expressible in the quantitative $\mu$-calculus [5, 18]. We claim that these metrics and relations represent the canonical extension of bisimulation to games. We also introduce asymmetrical versions of these metrics and equivalences, which constitute the canonical extension of simulation.

An equivalence relation for deterministic games that are either turn-based, or where the players are constrained to playing pure moves, has been introduced in [2] and called alternating bisimulation. Relations and metrics for the general case of concurrent games have so far proved elusive, with some previous attempts at their definition by a subset of the authors following a subtly flawed approach [4, 16]. The cause of the difficulty goes to the heart of the definition of bisimulation. In the definition of bisimulation for transition systems, for every state $s$, $t$ of bisimilar states, we require that if $s$ can go to a state $s'$, then $t$ should be able to go to $t'$, such that $s'$ and $t'$ are again bisimilar (as we also ask that $s$, $t$ have an equivalent predicate valuation). This definition has been extended to Markov decision processes by requiring that for every mixed move from $s$, there is a mixed move from $t$, such that the moves induce probability distributions over successor states that are equivalent modulo the underlying bisimulation [22, 21]. Unfortunately, the generalization of this appealing definition to games fails. It turns out, as we prove in this paper, that requiring players to be able to replicate probability distributions over successors (modulo the underlying equivalence) leads to an equivalence that is too fine, and that may fail to relate states at which the same quantitative $\mu$-calculus formulas hold. We show that phrasing the definition in terms of distributions over successor states is the wrong approach for games; rather, the definition should be phrased in terms of expectations of certain metric-bounded quantities.

While this introduction focused mostly on metrics and equivalence relations, we also develop results for the asymmetrical versions of these notions, related to simulation.
2. Games and Goals

We will develop metrics for game structures over a set \( S \) of states. We start with some preliminary definitions. For a finite set \( A \), let \( \text{Dist}(A) \) denote the set of probability distributions over \( A \). We say that \( p \in \text{Dist}(A) \) is deterministic if there is \( a \in A \) such that \( p(a) = 1 \).

For a set \( S \), a valuation over \( S \) is a function \( f : S \mapsto [0,1] \) associating with every element \( s \in S \) a value \( 0 \leq f(s) \leq 1 \); we let \( F \) be the set of all valuations. For \( c \in [0,1] \), we denote by \( c_{\in F} \) the constant valuation such that \( c_{\in F} = c \) at all \( s \in S \). We order valuations pointwise: for \( f,g \in F \), we write \( f \leq g \) iff \( f(s) \leq g(s) \) at all \( s \in S \); we remark that \( F \), under \( \leq \), forms a lattice.

Given \( a,b \in \mathbb{R} \), we write \( a \sqcup b = \max\{a,b\} \), and \( a \sqcap b = \min\{a,b\} \); we also let \( a \sqcup b = \min\{1,\max\{0,a+b\}\} \) and \( a \sqcap b = \max\{0,\min\{1, a-b\}\} \). We extend \( \sqcup, \sqcap \) to valuations by interpreting them in pointwise fashion.

A directed metric is a function \( d : S^2 \mapsto \mathbb{R}_{\geq 0} \) which satisfies \( d(s,s) = 0 \) and \( d(s,t) \leq d(s,u) + d(u,t) \) for all \( s,t,u \in S \). We denote by \( \mathcal{M} \subseteq S^2 \mapsto \mathbb{R} \) the space of all metrics; this space, ordered pointwise, forms a lattice which we indicate with \( (\mathcal{M}, \leq) \). Given a metric \( d \in \mathcal{M} \), we denote by \( \bar{d} \) its opposite version, defined by \( \bar{d}(s,t) = d(t,s) \) for all \( s,t \in S \); we say that \( d \) is symmetrical if \( d = \bar{d} \).

2.1. Game Structures

We assume a fixed, finite set \( V \) of observation variables. A (two-player, concurrent) game structure \( G = \langle S, [\cdot], \text{Moves}, \Gamma_1, \Gamma_2, \delta \rangle \) consists of the following components \([1, 3]:\)

- A finite set \( S \) of states.
- A variable interpretation \([\cdot] : V \times S \mapsto [0,1] \), which associates with each variable \( v \in V \) a valuation \([v]\).
- A finite set \( \text{Moves} \) of moves.
- Two move assignments \( \Gamma_1, \Gamma_2 : S \mapsto \text{Moves} \setminus \emptyset \). For \( i \in \{1,2\} \), the assignment \( \Gamma_i \) associates with each state \( s \in S \) the nonempty set \( \Gamma_i(s) \subseteq \text{Moves} \) of moves available to player-\( i \) at state \( s \).
- A probabilistic transition function \( \delta : S \times \text{Moves} \times \text{Moves} \mapsto \text{Dist}(S) \), that gives the probability \( \delta(s,a_1,a_2) \) of a transition from \( s \) to \( t \) when player-1 plays move \( a_1 \) and player-2 plays move \( a_2 \).

At every state \( s \in S \), player 1 chooses a move \( a_1 \in \Gamma_1(s) \), and simultaneously and independently player 2 chooses a move \( a_2 \in \Gamma_2(s) \). The game then proceeds to the successor state \( t \in S \) with probability \( \delta(s,a_1,a_2) \). We denote by \( \text{Dest}(s,a_1,a_2) = \{ t \in S \mid \delta(s,a_1,a_2)(t) > 0 \} \) the set of destination states when actions \( a_1,a_2 \) are chosen at \( s \). The variables in \( V \) naturally induce an equivalence on states: for states \( s,t \), define \( s \equiv t \) if for all \( v \in V \) we have \([v](s) = [v](t) \). In the following, unless otherwise noted, the definitions refer to a game structure with components \( G = \langle S, [\cdot], \text{Moves}, \Gamma_1, \Gamma_2, \delta \rangle \). For player \( i \in \{1,2\} \), we write \( \sim i = 3 - i \) for the opponent. We also consider the following subclasses of game structures.

- **Turn-based game structures.** A game structure \( G \) is turn-based if we can write \( S \) as the disjoint union of two sets: the set \( S_1 \) of player-1 states, and the set \( S_2 \) of player-2 states, such that \( s \in S_1 \) implies \( [\Gamma_2(s)] = 1 \), and \( s \in S_2 \) implies \( [\Gamma_1(s)] = 1 \), and further, there is a special variable turn \( \in V \), such that \( \text{turn}(s) = 1 \) iff \( s \in S_1 \), and \( \text{turn}(s) = 0 \) iff \( s \in S_2 \); thus, the variable turn indicates whose turn it is to play at a state. Turn-based games are often called perfect information games \([20] \).

- **Markov decision processes.** A game structure \( G \) is a Markov decision process (MDP) \([6] \) if only one of the two players has a choice of moves. For \( i \in \{1,2\} \), we say that a structure is an \( i \)-MDP if \( [\text{turn}] = 1 \) for all \( s \in S \). For MDPs, we omit the (single) move of the player without a choice of moves, and write \( \delta(s,a) \) for the transition function.

- **Deterministic game structures.** A game structure \( G \) is deterministic if, for all \( s \in S \), \( a_1 \in \text{Moves} \), and \( a_2 \in \text{Moves} \), there exists a \( t \in S \) such that \( \delta(s,a_1,a_2) \)(t) = 1 \); we denote such \( t \) by \( \tau(s,a_1,a_2) \). We sometimes call probabilistic a general game structure, to emphasize the fact that it is not necessarily deterministic.

**Pure and mixed moves.** A mixed move is a probability distribution over the moves available to a player at a state. We denote by \( D_i(s) = \text{Dist}(\Gamma_i(s)) \) the set of mixed moves available to player \( i \in \{1,2\} \) at \( s \in S \). The moves in \( \text{Moves} \) are called pure moves, in contrast to mixed moves. We extend the transition function to mixed moves. For \( s \in S \) and \( x_1 \in D_1(s) \), \( x_2 \in D_2(s) \), we write \( \delta(s,x_1,x_2) \) for the next-state probability distribution induced by the mixed moves \( x_1 \) and \( x_2 \), defined for all \( t \in S \) by

\[
\delta(s,x_1,x_2)(t) = \sum_{a_1 \in \Gamma_1(s)} \sum_{a_2 \in \Gamma_2(s)} \delta(s,a_1,a_2)(t) x_1(a_1) x_2(a_2).
\]

In the following, we sometimes restrict the moves of the players to pure moves. We identify a pure move \( a \in \Gamma_i(s) \) available to player \( i \in \{1,2\} \) at a state \( s \) with a deterministic distribution that plays \( a \) with probability 1.

**The deterministic setting.** The deterministic setting consists in considering deterministic game structures, and players restricted to play pure moves.
Predecessor operators. Given a valuation $f \in F$, a state $s \in S$, and two mixed moves $x_1 \in D_1(s)$ and $x_2 \in D_2(s)$, we define the expectation of $f$ from $s$ under $x_1, x_2$: $$E^{x_1, x_2}(f) = \sum_{t \in S} \delta(s, x_1, x_2)(t) f(t)$$

For a game structure $G$, for $i \in \{1, 2\}$ we define the valuation transformer $\text{Pre}_i : F \mapsto F$ by, for all $f \in F$ and $s \in S$, $$\text{Pre}_i(f)(s) = \sup_{x_i \in D_i(s)} \inf_{x_{\neg i} \in D_{\neg i}(s)} E^{x_1, x_2}(f).$$

Intuitively, $\text{Pre}_i(f)(s)$ is the maximal expectation player $i$ can achieve of $f$ after one step from $s$; this is the classical “one-day” or “next-stage” operator of the theory of repeated games [10]. We also define a deterministic version of this operator, in which players are forced to play pure moves: $$\text{Pre}_i^\Gamma(f)(s) = \max_{x_i \in \Gamma_i(s)} \min_{x_{\neg i} \in \Gamma_{\neg i}(s)} E^{x_1, x_2}(f).$$

2.2. Quantitative $\mu$-calculus

We consider the set of properties expressed by the quantitative $\mu$-calculus ($q\mu$). As discussed in [13, 5, 18], a large set of properties can be encoded in $q\mu$, spanning from basic properties such as maximal reachability and safety probability, to the maximal probability of satisfying a general $\omega$-regular specification.

Syntax. The syntax of quantitative $\mu$-calculus is defined with respect to the set of observation variables $V$ as well as a set $MVars$ of calculus variables, which are distinct from the observation variables in $V$. The syntax is given as follows:

$$\phi ::= c \mid v \mid Z \mid \neg \phi \mid \phi \lor \phi \mid \phi \land \phi \mid \phi \oplus c \mid \phi \ominus c \mid \text{pre}_i(\phi) \mid \text{pre}_2(\phi) \mid \mu Z. \phi \mid \nu Z. \phi$$

for constants $c \in [0, 1]$, observation variables $v \in V$, and calculus variables $Z \in MVars$. In the formulas $\mu Z. \phi$ and $\nu Z. \phi$, we furthermore require that all occurrences of the bound variable $Z$ in $\phi$ occur in the scope of an even number of occurrences of the complement operator $\neg$. A formula $\phi$ is closed if every calculus variable $Z$ in $\phi$ occurs in the scope of a quantifier $\mu Z$ or $\nu Z$. From now on, with abuse of notation, we denote by $q\mu$ the set of closed formulas of $q\mu$. A formula is a player-$i$ formula, for $i \in \{1, 2\}$, if $\phi$ does not contain the $\text{pre}_{\neg i}$ operator; we denote with $q\mu_i$ the syntactic subset of $q\mu$ consisting only of closed player-$i$ formulas. A formula is in positive form if the negation appears only in front of game variables, i.e., in the context $\neg v$; we denote with $q\mu^+$ and $q\mu_\neg$ the subsets of $q\mu$ and $q\mu_i$ consisting only of positive formulas.

We remark that the fixpoint operators $\mu$ and $\nu$ will not be needed to achieve our results on the logical characterization of game relations. They have been included in the calculus because they allow the expression of many interesting properties, such as safety, reachability, and in general, $\omega$-regular properties.

Semantics. A variable valuation $\xi : MVars \mapsto F$ is a function that maps every variable $Z \in MVars$ to a valuation in $F$. We write $[\xi][Z \mapsto f]$ for the valuation that agrees with $\xi$ on all variables, except that $Z$ is mapped to $f$. Given a game structure $G$ and a variable valuation $\xi$, every formula $\phi$ of the quantitative $\mu$-calculus defines a valuation $[\phi]^G_\xi \in F$ (the superscript $G$ is omitted if the game structure is clear from the context):

$$[c]_\xi = c \quad [v]_\xi = [v] \quad [Z]_\xi = \xi(Z)$$

$$[-\phi]_\xi = 1 - [\phi]_\xi \quad [\phi \lor c]_\xi = [\phi]_\xi [c]_\xi$$

$$[\phi_1 \land \phi_2]_\xi = [\phi_1]_\xi [\phi_2]_\xi$$

$$[\text{pre}_i(\phi)]_\xi = [\nu(Z)]_\xi [\phi]_\xi$$

$$[\mu(Z) \land \phi]_\xi = \{ [\inf_{s \in F} f \in F \mid f = [\phi]_\xi[Z \mapsto f] \} \}$$

where $i \in \{1, 2\}$. The existence of the fixpoints is guaranteed by the monotonicity and continuity of all operators and can be computed by Picard iteration [5]. If $\phi$ is closed, $[\phi]_\xi$ is independent of $\xi$, and we write simply $[\phi]$.

We also define a deterministic semantics $[\cdot]_\xi^\Gamma$ for $q\mu$, in which players can select only pure moves in the operators $\text{pre}_1, \text{pre}_2$. $[\cdot]_\xi^\Gamma$ is defined as $[\cdot]_\xi$, except for the clause

$$[\text{pre}_i(\phi)]_\xi^\Gamma = [\nu(Z)]_\xi [\text{pre}_i(\phi)]_\xi.$$

Example 1 Given a set $T \subseteq S$, the characteristic valuation $T$ of $T$ is defined by $T(s) = 1$ if $s \in T$, and $T(s) = 0$ otherwise. With this notation, the maximal probability with which player $i \in \{1, 2\}$ can ensure that $T \subseteq S$ is given by $[\nu Z.(T \land \text{pre}_i(Z))]$, and the maximal probability with which player $i$ can guarantee staying in $T$ forever is given by $[\nu Z.(T \land \text{pre}_i(Z))]$ [5].

3. Metrics

We are interested in developing a metric on states of a game structure that captures an approximate notion of equivalence: states close in the metric should yield similar values to the players for any winning objective. Specifically, we are interested in defining a bisimulation metric $[\approx]_g \in \mathcal{M}$ such that for any game structure $G$ and states $s, t$ of $G$, the following continuity property holds:

$$[\approx]_g(s, t) = \sup_{\phi \in q\mu} |[\phi]_\xi(s) - [\phi]_\xi(t)|.$$

In particular, the kernel of the metric, that is, states at distance 0, are equivalent: each player can get exactly the same value from either state for any objective. Notice that
in defining the metric independent of a player, we are ex-
we write
for any game, the value obtained by player 2 is 1 minus the
since the underlying games we consider are determined —
Thus, our metrics will generalize equivalence and re-
and the deterministic setting. To underline the connection be-
write [s ∼g t] for [∼g](s, t), so that the desired property
of the bisimulation metric can be stated as [s ∼g t] = supφ∈qφ[[φ](s)−[φ](t)]. Metrics of this type have already
been developed for Markov decision processes (MDPs)
[27, 8]. Our construction of metrics for games starts from
an analysis of these constructions.

3.1. Metrics for MDPs

We consider the case of 1-MDPs; the case for 2-MDPs
is symmetrical. Throughout this subsection, we fix a
1-MDP (S, [·], Moves, Γ1, Γ2, δ). Before we present the
metric correspondent of probabilistic simulation, we first
rephrase classical probabilistic (bi)simulation on MDPs
[15, 11, 22, 23] as a fixpoint of a relation transformer
post : 2S×S → 2S×S. For all relations R ⊆ S×S and
s, t ∈ S, let (s, t) ∈ F(R) iff

\[ s \equiv t \land \forall x_1 \in M_1(s) : \exists y_1 \in M_1(t) \land \delta(s, x_1) \subseteq_R \delta(t, y_1), \]

for all states s, t ∈ S. In (2), ≡ is the predicate equivalence
relation: s ≡ t if the predicates have the same value at s and
t. The relation ⊆_R lifts the relation R on states to a relation
on distributions. Precisely, for a relation R ⊆ S×S and
two distributions p, q ∈ Dist(S), we let p ⊆_R q if there is a
function Δ : S × S → [0, 1] such that (i) Δ(s, s') > 0 im-
plies (s, s') ∈ R, (ii) p(s) = \sum_{s' \in S} Δ(s, s') for any s ∈ S,
and (iii) q(s') = \sum_{s \in S} Δ(s, s') for any s' ∈ S. Probabilis-
similation is the greatest fixpoint of (2); probabilistic
bisimulation is the greatest symmetrical fixpoint of (2).

To obtain a metric equivalent of probabilistic simulation,
we lift the above fixpoint from relations (subsets of S^2) to
metrics (maps S^2 → R). First, we define [≡] ∈ M for all
s, t ∈ S by [s ≡ t] = max_{v \in V} [v](s) − [v](t). Second,
we lift (2) to metrics, defining a metric transformer H^{MDP}_{post} :
M → M. For all d ∈ M, let D(δ(s, x_1), δ(t, y_1))(d) be
the distribution distance between δ(s, x_1) and δ(t, y_1) with
respect to the metric d. We will show later how to define
such a distribution distance. For s, t ∈ S, we let

\[ H^{MDP}_{post}(d)(s, t) = [s \equiv t] \cup \sup_{x_1 \in \Gamma_1(s)} \inf_{y_1 \in \Gamma_1(t)} D(\delta(s, x_1), \delta(t, y_1))(d). \]

In this definition, the ∀ and ∃ of (2) have been replaced by
sup and inf, respectively. Since equivalent states should
have distance 0, the simulation metric in MDPs is defined as
the least (rather than greatest) fixpoint of (3) [27, 8]. Simi-
larly, the bisimulation metric is defined as the least symmet-
rical fixpoint of (3).

For a distance d ∈ M and two distributions p, q ∈ Dist(S),
the distribution distance D(p, q)(d) between p and q
with respect to d, is a measure of how much “work” we
have to do to make p look like q, given that moving a unit
of probability mass from s ∈ T to t ∈ S has cost d(s, t).
More precisely, D(p, q)(d) is defined via the trans-shipping
problem, as the minimum cost of shipping the distribution
p into q, with edge costs d. Thus, D(p, q)(d) is the solution
of the following linear programming (LP) problem over the
set of variables \{λ_{s,t}\}_{s,t \in S} [27]:

\[ \begin{align*}
\text{Minimize} & \quad \sum_{s,t \in S} d(s,t)\lambda_{s,t} \\
\text{subject to} & \quad \sum_{s \in S} \lambda_{s,t} = p(s), \\
& \quad \sum_{t \in S} \lambda_{s,t} = q(t), \\
& \quad \lambda_{s,t} \geq 0.
\end{align*} \]

Equivalently, we can define D(p, q)(d) via the dual of the
above LP problem. Given a metric d ∈ M, let C(d) ⊆ F
be the subset of valuations k ∈ F such that k(s) − k(t) ≤
d(s, t) for all s, t ∈ S. Then the dual formulation is:

\[ \begin{align*}
\text{Maximize} & \quad \sum_{s \in S} p(s)k(s) - \sum_{s \in S} q(s)k(s) \\
\text{subject to} & \quad k \in C(d).
\end{align*} \]

The constraint C(d) on the valuation k, states the value of k
across states cannot differ by more than d. This means, in-
tuitively, that k behaves like the valuation of a qu formula:
as we will see, the logical characterization implies that d is a
bound for the difference in valuation of qu formulas across
states. Indeed, the logical characterization of the metrics is
proved by constructing formulas whose valuation approxi-
mate that of the optimal k. Plugging (4) into (3), we obtain:

\[ H^{MDP}_{post}(d)(s, t) = [s \equiv t] \cup \sup_{x_1 \in \Gamma_1(s)} \inf_{y_1 \in \Gamma_1(t)} (\mathbb{E}^{\delta_1}_{s}(k) - \mathbb{E}^{\delta_1}_{t}(k)). \]

We can interpret this definition as follows. State t is try-
ing to simulate state s (this is a definition of a simulation met-
tric). First, state s chooses a mixed move x_1, attempting to
make simulation as hard as possible; then, state t chooses
a mixed move y_1, trying to match the effect of x_1. Once x_1
and y_1 have been chosen, the resulting distance between s
and t is equal to the maximal difference in expectation, for
moves x_1 and y_1, of a valuation k ∈ C(d). We call the met-
ric transformer H^{MDP}_{post} the a posteriori metric transformer:
the valuation k in (5) is chosen after the moves x_1 and y_1
are chosen. We can define an a priori metric transformer,
where \( k \) is chosen before \( x_1 \) and \( y_1 \):

\[
H_{\text{pro}}^{\text{MDP}}(d)(s,t) = [s \equiv t] \bigcup \sup_{k \in C(d)} \sup_{x_1 \in \Gamma_1(s)} \inf_{y_1 \in \Gamma_1(t)} \left( E_{x}^{s_1}(k) - E_{y}^{t_1}(k) \right)
\]

(6)

Intuitively, in the a priori transformer, first a valuation \( k \in C(d) \) is chosen. Then, state \( t \) must simulate state \( s \) with respect to the expectation of \( k \). State \( s \) chooses a move \( x_1 \), trying to maximize the difference in expectations, and state \( t \) chooses a move \( y_1 \), trying to minimize it. The distance between \( s \) and \( t \) is then equal to the difference in the resulting expectations of \( k \).

Theorem 1 below states that for MDPs, a priori and a posteriori simulation metrics coincide. In the next section, we will see that this is not the case for games.

**Theorem 1** For all MDPs, \( H_{\text{post}}^{\text{MDP}} = H_{\text{pro}}^{\text{MDP}} \).

**Proof** Consider two states \( s,t \in S \), and a metric \( d \in M \). We have to prove that \( \sup_{k \in C(d)} \sup_{x_1 \in \Gamma_1(s)} \inf_{y_1 \in \Gamma_1(t)} E_{x}^{s_1}(k) - E_{y}^{t_1}(k) = \sup_{k \in C(d)} \sup_{x_1 \in \Gamma_1(s)} \inf_{y_1 \in \Gamma_1(t)} E_{x}^{s_1}(k) - E_{y}^{t_1}(k) \). In the left-hand side, we can exchange the two outer sups. Then, noticing that the difference in expectation is bilinear in \( k \) and \( y_1 \) for a fixed \( x_1 \), that \( y_1 \) is a probability distribution, and that \( k \) is chosen from a compact convex subset, we apply the generalized maximin theorem [25] to exchange \( \sup_{k \in C(d)} \inf_{y_1 \in \Gamma_1(t)} \) into \( \inf_{y_1 \in \Gamma_1(t)} \sup_{k \in C(d)} \), thus obtaining the right-hand side. ■

The metrics defined above are logically characterized by \( q_\mu \). Precisely, let \([\sim] \in M \) be the least logical symmetrical fixpoint of \( H_{\text{pro}}^{\text{MDP}} = H_{\text{post}}^{\text{MDP}} \). Then, the results of [8] (originally stated for \( H_{\text{post}}^{\text{MDP}} \)) state that for all states \( s,t \) of a 1-MDP, we have \([s \sim t] = \sup_{k \in C(d)} [\phi(k)](s) - [\phi(k)](t) \).

**3.2. Metrics for Concurrent Games**

We now extend the simulation and bisimulation metrics from MDPs to general game structures. As we shall see, unlike for MDPs, the a priori and the a posteriori metrics do not coincide over games. In particular, we show that the a priori formulation satisfies both a tight logical characterization as well as reciprocity while, perhaps surprisingly, the more natural a posteriori version does not.

A posteriori metrics are defined via the metric transformer \( H_{\equiv_1} : \mathcal{M} \mapsto M \) as follows, for all \( d \in M \) and \( s,t \in S \):

\[
H_{\equiv_1}(d)(s,t) = [s \equiv t] \bigcup \sup_{x_1 \in D_1(s)} \inf_{y_1 \in D_1(t)} \left( E_{x}^{s_1}(k) - E_{y}^{t_1}(k) \right)
\]

(7)

We define an a posteriori game simulation metric \([\subseteq_1]\) as the least fixpoint of \( H_{\equiv_1} \), and we define an a posteriori game bisimulation metric \([\cong_1]\) as the least symmetrical fixpoint of \( H_{\equiv_1} \). The a posteriori simulation metric \([\subseteq_1]\) has been introduced in [4, 16].

A priori metrics are defined by bringing the \( \sup_{k \in C(d)} \) outside. Precisely, we define a metric transformer \( H_{\equiv_1} : \mathcal{M} \mapsto \mathcal{M} \) as follows, for all \( d \in M \) and \( s,t \in S \):

\[
H_{\equiv_1}(d)(s,t) = [s \equiv t] \bigcup \sup_{k \in C(d)} \inf_{x_1 \in D_1(s)} \sup_{y_1 \in D_1(t)} \left( E_{x}^{s_1}(k) - E_{y}^{t_1}(k) \right)
\]

(7)

The a priori simulation metric \([\subseteq_1]\) is the least fixpoint of \( H_{\equiv_1} \), and the a priori bisimulation metric \([\cong_1]\) is the least symmetrical fixpoint of \( H_{\equiv_1} \).

We now show some basic properties of these metrics. We show that a priori fixpoints give a (directed) metric. We show that a priori and a posteriori metrics are distinct. We then focus on the a priori metrics, and show through our results that they are the natural metrics for concurrent games.

**Theorem 2** For all game structures \( G \), and all states \( s,t,u \) of \( G \), we have (1) \([s \preceq_1 t] \geq 0 \), and \([s \preceq_1 t] \geq 0 \), and (2) \([s \preceq_1 u] \leq [s \preceq_1 t] + [t \preceq_1 u] \).

**Proof** The first part follows by considering \( k(s) = 0 \) for all \( s \in S \). The second part is proved by induction on the iterations used to compute the fixpoint of \( H_{\equiv_1} \), as given in (8). The crucial step consists in observing that, for \( d \in M \), we have

\[
\sup_{k \in C(d)} \left( \text{Pre}_1(k)(s) - \text{Pre}_1(k)(t) \right) \]

\[
+ \sup_{k \in C(d)} \left( \text{Pre}_1(k)(t) - \text{Pre}_1(k)(u) \right) \leq \sup_{k \in C(d)} \left( \text{Pre}_1(k)(s) - \text{Pre}_1(k)(u) \right). \]

A priori and a posteriori metrics are distinct. First, we show that a priori and a posteriori metrics are distinct in general: the a priori metric never exceeds the a posteriori one, and there are concurrent games where it is strictly smaller. Intuitively, this can be explained as follows. Simulation entails trying to simulate the expectation of a valuation \( k \), as we see from (7), (8). It is easier to simulate a state
$\delta (\ast, \ast, \ast)(w)$

$\delta (\ast, \ast, \ast)(w)$

Figure 1. Transition probabilities from states $s, t$ to states $u, w$.

$s$ from a state $t$ if the valuation is known in advance, as in a priori metrics (8), than if the valuation $k$ is chosen after all the moves have been chosen, as in a posteriori metrics (7).

As a special case, we shall see that equality holds for turn-based game structures, in addition to MDPs as we have seen in the previous subsection.

Theorem 3 The following assertions hold.

1. For all game structures $G$, and for all states $s, t$ of $G$, we have $[s \preceq_1 t] \leq [s \preceq t].$

2. There is a game structure $G$, and states $s, t$ of $G$, such that $[s \preceq_1 t] = 0$ and $[s \preceq t] > 0.$

3. For all turn-based game structures, we have $\leq_1 = \subseteq_1.$

Proof The first assertion is a consequence of the fact that, for all functions $f : \mathbb{R}^2 \rightarrow \mathbb{R}$, we have $\sup_x \inf_y f(x, y) \leq \inf_y \sup_x f(x, y)$ (intuitively, it is easier to maximize if we can choose $x$ after $y$). Repeated applications of this allows us to show that, for all $d \in \mathcal{M}$, we have $H_{\leq}(d) \leq H_{\subseteq}(d)$ (with pointwise ordering). The result then follows from the monotonicity of $H_{\leq}$ and $H_{\subseteq}$.

For the second assertion, we give an example where a priori distances are strictly less than a posteriori distances. Consider a game with states $S = \{s, t, u, w\}.$ States $u$ and $w$ are sink states with $[u \equiv w] = 1$; states $s$ and $t$ are such that $[s \equiv t] = 0.$ At states $s$ and $t$, player-2 has moves $\{f, g\}$. Player-1 has a single move $\{a\}$ at state $s$, and moves $\{b, c\}$ at state $t$. The moves from $s$ and $t$ lead to $u$ and $w$ with transition probabilities indicated in Figure 1. In the figure, the point $b, f$ indicates the probability of going to $u$ and $w$ when the move pair $(b, f)$ is played, with $\delta(s, b, f)(u) + \delta(s, b, f)(w) = 1$; similarly for the other move pairs. The thick line segment between the points $a, f$ and $a, g$ represents the transition probabilities arising when player 1 plays move $a$, and player 2 plays a mixed move (a mix of $f$ and $g$).

We show that, in this game, we have $[s \subseteq_1 t] > 0.$ Consider the metric $d$ where $d(u, w) = 1$ (recall that $[u \equiv w] = 1$, and note the other distances do not matter, since $u, w$ are the only two destinations). We need to show

$\forall y_1 \in \mathcal{D}_1(t). \forall y_2 \in \mathcal{D}_2(t). \forall x_2 \in \mathcal{D}_2(s). \exists k \in \mathcal{C}(d).$

$\left( E_k^{a,x_2}(k) - E_t^{y_1,y_2}(k) \right) > 0$ (9)

Consider any mixed move $y_1 = ab + (1 - \alpha)c$, where $b, c$ are the moves available to player 1 at $t$, and $0 \leq \alpha \leq 1$. If $\alpha < 1/3$, choose move $f$ from $t$ as $y_2$, and choose $k(u) = 1$, $k(w) = 0$. Otherwise, choose move $g$ from $t$ as $y_2$, and choose $k(u) = 0$, $k(w) = 1$. With these choices, the transition probability $\delta(t, y_1, y_2)$ will fall outside of the segment $[(a, f), (a, g)]$ in Figure 1. Thus, with the choice of $k$ above, we ensure that the difference in (9) is always positive.

To show that in the game we have $[s \preceq_1 t] = 0$, it suffices to show (given that $[s \preceq t] \geq 0$) that

$\forall k \in \mathcal{C}(d). \forall y_1 \in \mathcal{D}_1(t). \forall y_2 \in \mathcal{D}_2(t). \forall x_2 \in \mathcal{D}_2(s).$

$\left( E_k^{a,x_2}(k) - E_t^{y_1,y_2}(k) \right) \leq 0.$

If $k(u) = k(w)$, the result is immediate. Assume otherwise, without loss of generality, that $k(u) < k(w)$, and choose $y_1 = c$. For every $y_2$, the distribution of successor states (and of $k$-expectations) will be in the interval $[(c, f), (c, g)]$ in Figure 1. By choosing $x_2 = f$, we have that $E_k^{a,f}(k) < E_k^{c,f}(k)$ for all $y_2 \in \mathcal{D}_2(t)$, leading to the result.

The last assertion of the theorem is proven in the same way as Theorem 1. □

Reciprocity of a priori metric. The previous theorem establishes that the a priori and a posteriori metrics are in general distinct. We now prove that it is the a priori metric, rather than the a posteriori one, that enjoys reciprocity, and that provides a (quantitative) logical characterization of $[\mu_1].$ We begin by considering reciprocity.

Theorem 4 The following assertions hold.

1. For all game structures $G$, we have $[\preceq_1] = [\preceq_2], and [\subseteq_1] = [\subseteq_2].$

2. There is a concurrent game structure $G$, with states $s$ and $t$, where $[\subseteq_1] \neq [\subseteq_2].$

Proof For the first assertion, it suffices to show that, for all $d \in \mathcal{M}$, and states $s, t \in S$, we have $H_{\leq_1}(d)(s, t) = H_{\subseteq_1}(d)(t, s).$ We proceed as follows:

$\sup_{k \in \mathcal{C}(d)} (\text{Pre}_1(k)(s) - \text{Pre}_1(k)(t))$ (10)

$= \sup_{k \in \mathcal{C}(d)} (-\text{Pre}_2(1 - k)(s) + \text{Pre}_2(1 - k)(t))$ (11)

$= \sup_{k \in \mathcal{C}(d)} (\text{Pre}_2(k)(t) - \text{Pre}_2(k)(s))$ (12)

The step from (10) to (11) uses $\text{Pre}_1(k)(s) = 1 - \text{Pre}_2(1 - k)(s)$ [5], and the step from (11) to (12) uses the change of variables $k \rightarrow 1 - k.$
To show that reciprocity fails for a posteriori simulation metrics, consider the game from Figure 1. We add two new moves to player 2 at state $t$, namely $\{e, h\}$, such that for any $\alpha \in [0, 1]$, $\delta(t, s, \alpha e + (1 - \alpha)h)(u) = \delta(s, a, \alpha f + (1 - \alpha)g)(u)$: using move $\{e, h\}$ player 2 can ensure that transition probabilities are on the solid line in Figure 1. Reasoning as in Theorem 3, it still is the case that $s R t$.

As a consequence of this theorem, we write $\simeq_g$ in place of $\equiv_1 = \equiv_2$, to emphasize that the player-1 and player-2 versions of game equivalence metrics coincide.

**Logical characterization of a priori metric.** The following theorem expresses the fact that $[q\mu]_1$ provides a logical characterization for the a priori metrics. The proof of the theorem uses ideas from [16] and [8].

**Theorem 5**  The following assertions hold for all game structures $G$, and for all states $s, t$ of $G$:

$$
\begin{align*}
[s \preceq_1 t] &= \sup_{\phi \in q\mu_1} (|\phi(s) - \phi(t)|) \\
[s \simeq_2 t] &= \sup_{\phi \in q\mu_2} (|\phi(s) - \phi(t)|)
\end{align*}
$$

We note that, due to Theorem 3, an analogous result does not hold for the a posteriori metrics. Together with the lack of reciprocity of the a posteriori metrics, this is a strong indication that the a priori metrics, and not the a posteriori ones, are the “natural” metrics on concurrent games.

**The Kernel.** The kernel of the metric $\simeq_g$ defines an equivalence relation $\simeq_g$ on the states of a game structure: $s \simeq_g t$ iff $[s \simeq_g t] = 0$. We call this the game bisimulation relation. Notice that by the reciprocity property of $\simeq_g$, the game bisimulation relation is canonical: $\simeq_1 = \simeq_2 = \simeq_g$. Similarly, we define the game simulation preorder $s \preceq_1 t$ as the kernel of the directed metric $[\preceq_1]$, that is, $s \preceq_1 t$ iff $[s \preceq_1 t] = 0$. Alternatively, it is possible to define $\preceq_1$ and $\preceq_g$ directly. Given a relation $R \subseteq S \times S$, let $B(R) \subseteq F$ consist of all valuations $k \in F$ such that, for all $s, t \in S$, if $sRt$ then $k(s) \leq k(t)$. We have the following result.

**Theorem 6**  Given a game structure $G$, $\preceq_1$ (resp. $\simeq_1$) can be characterized as the largest (resp. largest symmetrical) relation $R$ such that, for all states $s, t$ with $sRt$, we have $s \equiv t$ and

$$
\forall k \in B(R) \forall x_1 \in D_1(s) \exists y_1 \in D_1(t) \forall y_2 \in D_2(t).
\exists x_2 \in D_2(s), (\mathbb{E}_{s,s_1}^{x_1, y_1}(k) \geq \mathbb{E}_{s,s_1}^{x_2, y_2}(k)).
$$

We note that the above theorem allows the computation of $\preceq_1$ via a partition-refinement scheme. From the logical characterization theorem, we obtain the following corollary.

**Corollary 1**  For any game structure $G$ and states $s, t$ of $G$, we have $s \simeq_g t$ iff $[\phi](s) = [\phi](t)$ holds for every $\phi \in q\mu$ and $s \preceq_1 t$ iff $[\phi](s) \leq [\phi](t)$ holds for every $\phi \in q\mu_1^+$.

**Computation.** The next theorem states that the metrics are computable to any degree of precision. This follows, since the definition of the distance between two states of a given game, as the least fixpoint of the metric transformer (8), can be written as a formula in the theory of reals, which is decidable [26]. Since the distance between two states may not be rational, we can only guarantee an approximate computation in general.

**Theorem 7**  For any game structure $G$, and states $s, t$ of $G$, the following assertions hold.

1. For all rational $v$, and all $\epsilon > 0$, it is decidable if $|[s \preceq_1 t] - v| < \epsilon$ and if $|[s \simeq_g t] - v| < \epsilon$.
2. It is decidable if $s \preceq_1 t$ and if $s \simeq_g t$.

**Game Metrics and (Bi-)simulation Metrics.** The a priori metrics assume an adversarial relationship between the players. We show that, on turn-based games, the a priori bisimulation metric coincides with the classical bisimulation metric where the players cooperate. Moreover, on 1-MDPs, the player-1 a priori simulation metric coincides with the cooperative simulation metric. The metric analog of classical (bi)simulation [19, 22] is obtained through the metric transformers $H_{\preceq_1} : M \mapsto M$ and $H_{\simeq_1} : M \mapsto M$ given by

$$
H_{\preceq_1}(d)(s, t) = [s \equiv t] \cup \sup_k \inf_k \sup_{k \in C(d)} \sup_{x_1 \in D_1(s)} \inf_{y_1 \in D_1(t)} \inf_{y_2 \in D_1(t)} \sup_{x_2 \in D_2(s)} \inf_{x_2 \in D_2(s)} \mathbb{E}_{x_1, y_1}^{x_2, y_2}(k).
$$

The metrics $[\preceq_1]$ and $[\simeq_1]$ are defined as the least fixed points of $H_{\preceq_1}$ and $H_{\simeq_1}$ respectively. The kernel of these metrics define the classical probabilistic simulation and bisimulation relations.

**Theorem 8**  The following assertions hold.

1. On turn-based game structures, $[\simeq_g] = [\simeq_1]$.
2. There is a deterministic game structure $G$ and states $s, t$ in $G$ such that $|s \simeq_g t| > |s \simeq_1 t|$.
3. There is a deterministic game structure $G$ and states $s, t$ in $G$ such that $|s \simeq_g t| < |s \simeq_1 t|$.

The first part follows easily from the definition. The second part is proven by the game in Figure 2, where it holds that $|s \simeq_g t| = \frac{1}{2}$ and $|s \simeq_1 t| = 0$. In this figure, as in all subsequent ones, different state colors denote that observation variables have different values at the states, so that the
4. Discussion

Our derivation of \(\simeq_i\) and \(\simeq_g\) for \(i \in \{1, 2\}\), as kernels of metrics, seems somewhat abstruse: most equivalence or similarity relations have been defined, after all, without resorting to metrics. We now point out how a generalization of the usual definitions [22, 2, 7, 8], suggested in [4, 16], fails to produce the “right” relations. Furthermore, the flawed relations obtained as a generalization of [22, 2, 7, 8] are no simpler than our definitions, based on kernel metrics. Thus, our study of game relations as kernels of metrics carries no drawbacks in terms of leading to more complicated definitions. Indeed, we believe that the metric approach is the superior one for the study of game relations.

We outline the flawed generalization of [22, 2, 7, 8] as proposed in [4, 16], explaining why it would seem a natural generalization. The alternating simulation of [2] is defined over deterministic game structures. Player-\(i\) alternating simulation, for \(i \in \{1, 2\}\), is the largest relation \(R\) satisfying the following conditions, for all states \(s, t \in S\): \(sRt\) implies \(s \equiv t\) and \(\forall a_i \in \Gamma_i(s) . \exists y_i \in \Gamma_i(t) . \forall y_{-i} \in \Gamma_{-i}(t) . \exists x_{-i} \in \Gamma_{-i}(s) . \tau(s, x_1, x_2) R \tau(t, y_1, y_2)\).

The MDP relations of [22], later extended to metrics by [7, 8], rely on the fixpoint (2), where sup plays the role of \(\forall\), inf plays the role of \(\exists\), and \(R\) is replaced by distribution equality modulo \(\sqsubseteq\), or \(\sqsubseteq\). This strongly suggests — incorrectly — that equivalences for general games (probabilistic, concurrent games) can be obtained by taking the double alternating of \(\forall \exists \forall \exists\) in the definition of alternating simulation, changing all \(\forall\) into sup, all \(\exists\) into inf, and replacing \(R\) by \(\sqsubseteq\). The definition that would result is as follows. We parameterize the new relations by a player \(i \in \{1, 2\}\), as well as by whether mixed moves or only pure moves are allowed. For a relation \(R \subseteq S \times S\), for \(M \in \{\Gamma, D\}\), for all \(s, t \in S\) and \(i \in \{1, 2\}\) consider the following conditions:

- (loc) \(sRt\) implies \(s \equiv t\).
- (M-i-alsim) \(sRt\) implies \(\forall x_i \in M_i(s) . \exists y_i \in M_i(t) . \forall y_{-i} \in M_{-i}(t) . \exists x_{-i} \in M_{-i}(s) . \delta(s, x_1, x_2) R \delta(t, y_1, y_2)\).

We then define the following relations:

- For \(i \in \{1, 2\}\) and \(M \in \{\Gamma, D\}\), player-\(i\) \(M\)-alternating simulation \(\sqsubseteq^M_i\) is the largest relation that satisfies (loc) and (M-i-alsim).
- For \(i \in \{1, 2\}\) and \(M \in \{\Gamma, D\}\), player-\(i\) \(M\)-alternating bisimulation \(\equiv^M_i\) is the largest symmetrical relation that satisfies (loc) and (M-i-alsim).

Over deterministic game structures, the definitions of \(\sqsubseteq^\Gamma_i\) and \(\equiv^\Gamma_i\) coincide with the alternating simulation and bisimulation relations of [2]. In fact, \(\sqsubseteq^\Gamma_i\) and \(\equiv^\Gamma_i\) capture the deterministic semantics of \(q\mu\), and thus in some sense generalize the results of [2] to probabilistic game structures.

Theorem 10 For any game structure \(G\) and states \(s, t\) of \(G\), the following assertions hold:

1. \(s \sqsubseteq^\Gamma_i t\) iff \([\phi]^G(s) = [\phi]^G(t)\) holds for every \(\phi \in q\mu_i\).
2. \(s \equiv^\Gamma_i t\) iff \([\phi]^G(s) \leq [\phi]^G(t)\) holds for every \(\phi \in q\mu_i^+\).

The following lemma states that \(\sqsubseteq^P_i\) and \(\equiv^P_i\) are the kernels of \([\equiv_i]\) and \([\sim_i]\), connecting thus the result of combining the definitions of [22] and [2] with a posteriori metrics.

Lemma 1 For all game structures \(G\), all players \(i \in \{1, 2\}\), and all states \(s, t\) of \(G\), we have \(s \sqsubseteq^P_i t\) iff \([s \sqsubseteq_i t] = 0\), and \(s \equiv^P_i t\) iff \([s \sim_i t] = 0\).
We are now in a position to prove that neither the $\Gamma$-relations not the $D$-relations are the “correct” relations on general concurrent games, since neither characterizes $[\mu i]$. In particular, the $D$-relations are too fine, and the $\Gamma$-relations are incomparable with the relations $\preceq_i$ and $\simeq_g$, for $i \in \{1, 2\}$. We prove these negative results first for the $D$-relations. They follow from Theorem 3 and 5.

**Theorem 11** The following assertions hold:

1. For all game structures $G$, all states $s, t$ of $G$, and all $i \in \{1, 2\}$, we have that $s \sqsubseteq_i^D t$ implies $s \preceq_i t$, and $s \not\sqsubseteq_i^D t$ implies $s \simeq_i t$.
2. There is a game structure $G$, and states $s, t$ of $G$, such that $s \preceq_i t$ but $s \not\sqsubseteq_i^D t$.
3. There is a game structure $G$, and states $s, t$ of $G$, such that $[\phi](s) = [\phi](t)$ for all $\phi \in [\mu i]$, but $s \not\sqsubseteq_i^D t$ for some $i \in \{1, 2\}$.

We now turn our attention to the $\Gamma$-relations, showing that they are incomparable with $\preceq_i$ and $\simeq_g$, for $i \in \{1, 2\}$.

**Theorem 12** The following assertions hold:

1. There exists a deterministic game structure $G$ and states $s, t$ of $G$ such that $s \sqsubseteq_1^\Gamma t$ but $s \not\simeq_1 t$, and $s \not\simeq_1^\Gamma t$ but $s \not\preceq_1 g t$.
2. There exists a turn-based game structure $G$ and states $s, t$ of $G$ such that $s \simeq_1 t$ but $s \not\simeq_1 t$, and $s \not\simeq_1^\Gamma t$.

Finally, we remark that, in view of Theorem 6, the definitions of the relations $\preceq_i$ and $\simeq_g$ for $i \in \{1, 2\}$ are no more complex than the definitions of $\sqsubseteq_i^D$, $\sqsubseteq_1^\Gamma$, $\equiv_i^D$, and $\equiv_1^\Gamma$.

**References**